Systems Dynamics

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Lecture 10 Solution of the Prediction Problem

10. Solution of the Prediction Problem

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Solution of the Prediction Problem

Consider the process v(t) with rational complex spectrum:

$$\underbrace{\xi(t)}_{\widehat{W}(z)} \underbrace{v(t)}_{v(t)}$$

where $\xi(\cdot) \sim WN(0, \lambda^2)$ and $\widehat{W}(z) = \frac{N(z)}{D(z)}$ is the spectral canonical factor, that is:

- + N(z) and D(z) are monic, co-prime and of the same degree
- All roots of N(z) (zeros of $\widehat{W}(z)$) have $|\cdot| \leq 1$
- All roots of D(z) (poles of $\widehat{W}(z)$) have $|\cdot| < 1$

• Then:

$$v(t) = \sum_{i=0}^{\infty} \hat{w}(i)\,\xi(t-i) = \hat{w}(0)\xi(t) + \hat{w}(1)\xi(t-1) + \cdots$$

where $\hat{w}(0), \hat{w}(1), \ldots$ are the samples of the impulse response of the system with transfer function $\widehat{W}(z)$:

$$\hat{w}(k) = \mathcal{Z}^{-1}\left[\widehat{W}(z)\right]$$

• Let us introduce the **additional assumption**:

All roots of N(z) (zeros of $\widehat{W}(z)$) have $|\cdot|<1$ Hence: the spectral factorisation theorem also holds for

$$\widetilde{W}(z) = \frac{1}{\widehat{W}(z)}$$

• Then, we are able to consider

$$v(t)$$
 $\widetilde{W}(z)$ $\xi(t)$

that is, feeding the system having transfer function $\widetilde{W}(z)$ with the process v(t), at the output we obtain exactly the white process $\xi(t)$

+ $\widetilde{W}(z)$ is called whitening filter and

$$\xi(t) = \sum_{i=0}^{\infty} \tilde{w}(i) v(t-i) = \tilde{w}(0)v(t) + \tilde{w}(1)v(t-1) + \cdots$$

• Thus, the whitening filter is kind of a **inverse filter** with respect to the canonical representation of the original process v(t)

· Let us now consider

$$v(t) = \sum_{i=0}^{\infty} \hat{w}(i)\,\xi(t-i) = \hat{w}(0)\xi(t) + \hat{w}(1)\xi(t-1) + \cdots$$

Clearly $v(t) \in \mathcal{H}_t[\xi]$ where we recall that $\mathcal{H}_t[\xi]$ is the space of all infinite linear combinations of $\xi(t), \xi(t-1), \ldots$. Analogously:

$$v(t-1) = \sum_{\substack{i=0\\\infty}}^{\infty} \hat{w}(i)\,\xi(t-1-i) = \hat{w}(0)\xi(t-1) + \hat{w}(1)\xi(t-2) + \cdots$$
$$v(t-2) = \sum_{\substack{i=0\\i=0}}^{\infty} \hat{w}(i)\,\xi(t-2-i) = \hat{w}(0)\xi(t-2) + \hat{w}(1)\xi(t-3) + \cdots$$

• Hence linear combinations of v(t), v(t-1), ... can be expressed as linear combinations of $\xi(t)$, $\xi(t-1)$, ... which implies:

$$\mathcal{H}_t[v] \subseteq \mathcal{H}_t[\xi]$$

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In the same way, one gets:

$$\xi(t) = \sum_{i=0}^{\infty} \tilde{w}(i) v(t-i) = \tilde{w}(0)v(t) + \tilde{w}(1)v(t-1) + \cdots$$

Clearly $\xi(t) \in \mathcal{H}_t[v]$ and

$$\xi(t-1) = \sum_{\substack{i=0\\\infty}}^{\infty} \tilde{w}(i) v(t-1-i) = \tilde{w}(0)v(t-1) + \tilde{w}(1)v(t-2) + \cdots$$

$$\xi(t-2) = \sum_{\substack{i=0\\\infty}}^{\infty} \tilde{w}(i) v(t-2-i) = \tilde{w}(0)v(t-2) + \tilde{w}(1)v(t-3) + \cdots$$

...

• Hence linear combinations of $\xi(t), \xi(t-1), \ldots$ can be expressed as linear combinations of $v(t), v(t-1), \ldots$ which implies:

$$\mathcal{H}_t[\xi] \subseteq \mathcal{H}_t[v]$$

• Thus, we finally conclude that:

$$\mathcal{H}_t[\xi] = \mathcal{H}_t[v]$$

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The Prediction Problem

- Given the rational spectrum stationary process v(t), we want to estimate $v(t+r), r \ge 1$ as a function of the past observations $v(t), v(t-1), \ldots$
- The observations $v(t), v(t-1), \ldots$ clearly make up an **a-priori knowledge** with respect to the quantity to be estimated v(t+r)
- Therefore, it is quite natural to cast the prediction problem in the framework of **Bayes estimation**:

$$\hat{v}(t+r \mid t) = E[v(t+r) \mid v(t), v(t-1), \ldots]$$

• Recall the geometric interpretation of the Bayes estimation:

$$\hat{\vartheta} = \frac{\lambda_{\vartheta d}}{\lambda_{dd}} d = \|\vartheta\| \cos \alpha \frac{d}{\|d\|}$$



Hence, in the case of the prediction problem $\hat{v}(t+r | t)$ is the projection of v(t+r) (interpreted as a geometric vector) on the subspace (hyper-plane) $\mathcal{H}_t[\xi] (= \mathcal{H}_t[v])$



• Let us now determine $\hat{v}(t+r \,|\, t)$:

$$\begin{aligned} v(t+r) &= \sum_{i=0}^{\infty} \hat{w}(i) \,\xi(t+r-i) \\ &= \underbrace{\hat{w}(0)\xi(t+r) + \hat{w}(1)\xi(t+r-1) + \dots + \hat{w}(r-1)\xi(t+1)}_{\alpha(t)} \\ &+ \underbrace{\hat{w}(r)\xi(t) + \hat{w}(r+1)\xi(t-1) + \dots}_{\beta(t)} \\ &= \alpha(t) + \beta(t) \end{aligned}$$

where:

- + $\alpha(t)$: lin. comb. of white process samples in $[t+1,t+r] \cap \mathbb{Z}$
- $\beta(t)$: lin. comb. of white process samples in $(-\infty, t] \cap \mathbb{Z}$



- But: $\xi(t)$ is white $\implies \alpha(t)$ and $\beta(t)$ are uncorrelated
- Hence, vectors associated with $\alpha(t)$ and $\beta(t)$ are orthogonal



• Thus: the **optimal prediction** coincides with $\beta(t)$:

$$\hat{v}(t+r \,|\, t) = \hat{w}(r)\xi(t) + \hat{w}(r+1)\xi(t-1) + \cdots$$

• Instead, the **prediction error** coincides with $\alpha(t)$ which is orthogonal to $\mathcal{H}_t[\xi] (= \mathcal{H}_t[v])$:

$$\varepsilon(t) = v(t+r) - \hat{v}(t+r \mid t) = \hat{w}(0)\xi(t+r) + \hat{w}(1)\xi(t+r-1) + \dots + \hat{w}(r-1)\xi(t+1)$$

Therefore, by defining

$$\widehat{W}_r(z) = \widehat{w}(r) + \widehat{w}(r+1)z^{-1} + \cdots$$

Optimal Predictor

 $\widehat{W}_r(z)$ is the transfer function of the *r*-th steps ahead optimal predictor from the white process samples $\xi(t)$

$$\underbrace{\xi(t)}_{\widehat{W}_r(z)} \qquad \underbrace{\hat{v}(t+r \mid t)}_{i}$$

Solution of the Prediction Problem

Determination of the Predictor

Determination of the Predictor

The computation of $\widehat{W}_r(z)$ is very simple: just carry out the **long-division** between the numerator and denominator of $\widehat{W}(z)$:

$$\begin{aligned} \widehat{W}(z) &= \widehat{w}(0) + \widehat{w}(1)z^{-1} + \dots + \widehat{w}(r-1)z^{-r+1} \\ &+ \widehat{w}(r)z^{-r} + \widehat{w}(r+1)z^{-r-1} + \dots \\ &= \widehat{w}(0) + \widehat{w}(1)z^{-1} + \dots + \widehat{w}(r-1)z^{-r+1} \\ &+ z^{-r} \left[\widehat{w}(r) + \widehat{w}(r+1)z^{-1} + \dots \right] \\ &= \widehat{w}(0) + \widehat{w}(1)z^{-1} + \dots + \widehat{w}(r-1)z^{-r+1} + z^{-r} \widehat{W}_r(z) \end{aligned}$$

Determination of the Optimal Predictor

 $\widehat{W}_r(z)$ is obtained as a result of the *r*-times repeated division: the **remainder**, multiplied by z^r is the $\widehat{W}_r(z)$ we were looking for:

$$\widehat{W}(z) = \frac{N(z)}{D(z)} \implies \frac{N(z)}{D(z)} = E(z) + z^{-r}\widehat{W}_r(z)$$

Consider:

$$\begin{aligned} v(t) + \frac{5}{6}v(t-1) + \frac{1}{6}v(t-2) &= \xi(t) + \frac{1}{9}\xi(t-1) \\ \implies (1 + \frac{5}{6}z^{-1} + \frac{1}{6}z^{-2})v(t) &= (1 + \frac{1}{9}z^{-1})\xi(t) \\ \implies v(t) &= \frac{z(z + \frac{1}{9})}{z^2 + \frac{5}{6}z + \frac{1}{6}}\xi(t) \end{aligned}$$

The assumptions of the spectral factorization theorem are satisfied because the poles are $-\frac{1}{2}, -\frac{1}{3}$ and the zeros are $0, -\frac{1}{9}$ and hence they lie strictly inside the unit-circle.

One-step ahead predictor:

$$\widehat{W}(z) = \mathbf{1} + z^{-1} \frac{-\frac{13}{18} - \frac{1}{6}z^{-1}}{1 + \frac{5}{6}z^{-1} + \frac{1}{6}z^{-2}} \Longrightarrow \widehat{W}_1(z) = \frac{-\frac{13}{18} - \frac{1}{6}z^{-1}}{1 + \frac{5}{6}z^{-1} + \frac{1}{6}z^{-2}}$$

Hence:

$$\hat{v}(t+1|t) = -\frac{5}{6}\hat{v}(t|t-1) - \frac{1}{6}\hat{v}(t-1|t-2) - \frac{13}{18}\xi(t) - \frac{1}{6}\xi(t-1)$$

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Determination of the Predictor: Basic Example (cont.)

Two-steps ahead predictor:

$$\begin{array}{c} 1 + \frac{1}{9}z^{-1} \\ 1 + \frac{5}{6}z^{-1} + \frac{1}{6}z^{-2} \\ \hline -\frac{13}{18}z^{-1} - \frac{1}{6}z^{-2} \\ -\frac{\frac{13}{18}z^{-1} - \frac{65}{108}z^{-2} - \frac{13}{108}z^{-3}}{\frac{47}{108}z^{-2} + \frac{13}{108}z^{-3}} \end{array} \right|$$

$$\widehat{W}(z) = 1 - \frac{13}{18}z^{-1} + z^{-2} \frac{\frac{47}{108} + \frac{13}{108}z^{-1}}{1 + \frac{5}{5}z^{-1} + \frac{1}{6}z^{-2}} \Longrightarrow \widehat{W}_2(z) = \frac{\frac{47}{108} + \frac{13}{108}z^{-1}}{1 + \frac{5}{5}z^{-1} + \frac{1}{6}z^{-2}}$$

Hence:

$$\hat{v}(t+2|t) = -\frac{5}{6}\hat{v}(t+1|t-1) - \frac{1}{6}\hat{v}(t|t-2) + \frac{47}{108}\xi(t) + \frac{13}{108}\xi(t-1)$$

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Determination of the Predictor from Observed Data

• Starting from the spectral canonical factor $\widehat{W}(z) = \frac{C(z)}{A(z)}$ we have obtained the transfer function $\widehat{W}_r(z) = \frac{C_r(z)}{A(z)}$ of the *r*-th steps ahead optimal predictor from the samples of $\xi(t)$:

$$\underbrace{\xi(t)}_{\widehat{W}_r(z)} \qquad \underbrace{\hat{v}(t+r|t)}_{\widehat{v}(t+r|t)}$$

• However, the process $\xi(t)$ is just a **mathematical abstraction** but certainly it is not a measurable entity. Instead, the goal is to determine a predictor yielding the prediction $\hat{v}(t+r|t)$ using the **measurable** past observations $v(t), v(t-1), v(t-2), \ldots$

Determination of the Predictor from Observed Data

• Recall that $\mathcal{H}_t[\xi] = \mathcal{H}_t[v]$. Hence, it is sufficient to suitably use the **whitening filter**:

$$\underbrace{\frac{v(t)}{\widetilde{W}(z)}}_{W_r(z)} \underbrace{\xi(t)}_{\widehat{W}_r(z)} \underbrace{\hat{v}(t+r|t)}_{W_r(z)}$$

$$W_r(z)$$

$$\widetilde{W}(z) = \frac{1}{\widehat{W}(z)} = \frac{A(z)}{C(z)} \implies W_r(z) = \frac{A(z)}{C(z)} \frac{C_r(z)}{A(z)} = \frac{C_r(z)}{C(z)}$$

Remark. The additional assumption for which the zeroes of C(z) should lie strictly inside the unit-circle is unavoidable to **guarantee the stability of the predictor**.

Determination of the Predictor: Basic Example (cont.)

Continuing the previous example, the one-step ahead predictor from the observed data is:

$$W_1(z) = \frac{1 + \frac{5}{6}z^{-1} + \frac{1}{6}z^{-2}}{1 + \frac{1}{9}z^{-1}} \cdot \frac{-\frac{13}{18} - \frac{1}{6}z^{-1}}{1 + \frac{5}{6}z^{-1} + \frac{1}{6}z^{-2}} = \frac{-\frac{13}{18} - \frac{1}{6}z^{-1}}{1 + \frac{1}{9}z^{-1}}$$

and hence

$$\hat{v}(t+1|t) = -\frac{1}{9}\hat{v}(t|t-1) - \frac{13}{18}v(t) - \frac{1}{6}v(t-1)$$

Analogously, the two-steps ahead predictor from the observed data is:

$$W_2(z) = \frac{1 + \frac{5}{6}z^{-1} + \frac{1}{6}z^{-2}}{1 + \frac{1}{9}z^{-1}} \cdot \frac{\frac{47}{108} + \frac{13}{108}z^{-1}}{1 + \frac{5}{6}z^{-1} + \frac{1}{6}z^{-2}} = \frac{\frac{47}{108} + \frac{13}{108}z^{-1}}{1 + \frac{1}{9}z^{-1}}$$

and hence

$$\hat{v}(t+2|t) = -\frac{1}{9}\hat{v}(t+1|t-1) + \frac{47}{108}v(t) + \frac{13}{108}v(t-1)$$

Solution of the Prediction Problem

Prediction Errors

Prediction Errors

One has:

$$\varepsilon(t) = v(t+r) - \hat{v}(t+r \mid t) = \hat{w}(0)\xi(t+r) + \hat{w}(1)\xi(t+r-1) + \dots + \hat{w}(r-1)\xi(t+1)$$

Notice that $\varepsilon(t)$ is a MA(r) process. Therefore:

•
$$E[\varepsilon(t)] = \hat{w}(0)E[\xi(t+r)] + \hat{w}(1)E[\xi(t+r-1)] + \dots + \hat{w}(r-1)E[\xi(t+1)] = 0$$

•
$$\operatorname{var}[\varepsilon(t)] = \left[\hat{w}(0)^2 + \hat{w}(1)^2 + \dots + \hat{w}(r-1)^2\right] \lambda^2$$

Remark. The variance of the prediction error **increases** as r **increases** and asymptotically converges to the variance of the process v(t) (the variance is finite thanks to the stability assumption).

Solution of the Prediction Problem

A Key Example

We want to solve the prediction problems for a generic AR(1) process.

$$v(t) = av(t-1) + \xi(t), \quad \xi(\cdot) \sim WN(0, \lambda^2), \quad |a| < 1$$

Hence:

$$(1 - az^{-1}) v(t) = \xi(t) \implies v(t) = \frac{1}{1 - az^{-1}}\xi(t) = \frac{1}{A(z)}\xi(t)$$

Since |a| < 1 , it follows that $\widehat{W}(z) = \frac{1}{A(z)}$ is a canonical factor.

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A notable/Key Example (cont.)

Then:

$$\widehat{W}(z) = \frac{1}{A(z)} = \frac{z}{z-a} = 1 + az^{-1} + a^2 z^{-2} + \dots + z^{-r} \frac{a^r z}{z-a}$$

Hence:

$$\widehat{W}_r(z) = \frac{a^r z}{z - a} = \frac{a^r}{1 - az^{-1}} \Longrightarrow \widehat{v}(t + r \,|\, t) = a\widehat{v}(t + r - 1 \,|\, t - 1) + a^r \,\xi(t)$$

$$W_r(z) = \frac{C_r(z)}{C(z)} = \frac{a^r z}{z} = a^r \Longrightarrow \hat{v}(t+r | t) = a^r v(t)$$

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A notable/Key Example (cont.)

- The outcome for which $\hat{v}(t+r|t) = a^r v(t)$ is not surprising: we have the process $v(t) = av(t-1) + \xi(t)$ and hence it is reasonable that the one-step ahead prediction of v(t+1) is av(t) as, at time t, a white noise is added to v(t).
- Notice that $\hat{v}(t+r|t) = a^r v(t) \longrightarrow 0$ for $r \to \infty$. This is consistent with E[v(t)] = 0 and then, for $r \to \infty$, the prediction has to coincide with the expected value of the process
- Prediction error variance:

$$\begin{aligned} \varepsilon(t) &= v(t+r) - \hat{v}(t+r \,|\, t) \\ &= a^r v(t) + \xi(t+r) + a\xi(t+r-1) + \dots + a^{r-1}\xi(t+1) - a^r v(t) \\ &= \xi(t+r) + a\xi(t+r-1) + \dots + a^{r-1}\xi(t+1) \end{aligned}$$

A notable/Key Example (cont.)

• Therefore, the prediction error is a MA(r-1) process for which

var
$$[\varepsilon(t)] = \left[1 + a^2 + a^4 + \dots + a^{2(r-1)}\right] \lambda^2$$

and hence the variance of the prediction error grows with respect to $\ r$.

Moreover:

$$\lim_{r \to \infty} \operatorname{var} \left[\varepsilon(t) \right] = \frac{\lambda^2}{1 - a^2} = \operatorname{var} \left[v(t) \right]$$

because

$$\operatorname{var}[v(t)] = E\left[[v(t)^{2}] = E\left\{\left[\sum_{i=0}^{\infty} \hat{w}(i)\,\xi(t-i)\right]^{2}\right\}\right\}$$
$$= \sum_{i=0}^{\infty} \hat{w}(i)^{2} E\left[\xi(t-i)^{2}\right] = \lambda^{2} \sum_{i=0}^{\infty} \hat{w}(i)^{2} = \lambda^{2} \frac{1}{1-a^{2}}$$

Solution of the Prediction Problem

One-step Ahead Prediction for ARMA Processes

One-step Ahead Prediction for ARMA Processes

• Consider the process $ARMA(n_a, n_c), \, \xi(\cdot) \sim WN(0, \lambda^2)$:

$$v(t) = a_1 v(t-1) + a_2 v(t-2) + \dots + a_n v(t-n) +\xi(t) + c_1 \xi(t-1) + c_2 \xi(t-2) + \dots + c_n \xi(t-n)$$

Hence:

$$A(z)v(t) = C(z)\xi(t)$$

with

$$A(z) = 1 - a_1 z^{-1} - \dots - a_{n_a} z^{-n_a}$$

$$C(z) = 1 + c_1 z^{-1} + \dots + c_{n_c} z^{-n_c}$$

$$\underbrace{\overset{\xi(t)}{\underbrace{W(z)}}_{w(z)} v(t)}_{v(z)} W(z) = \frac{C(z)}{A(z)} = \frac{1 + c_1 z^{-1} + \dots + c_{n_c} z^{-n_c}}{1 - a_1 z^{-1} - \dots - a_{n_a} z^{-n_a}}$$

Setting $n = \max(n_a, n_c)$: $W(z) = \frac{z^n + c_1 z^{n-1} + \dots + c_{n_c} z^{n-n_c}}{z^n - a_1 z^{n-1} - \dots - a_{n_a} z^{n-n_a}}$

One-step Ahead Prediction for ARMA Processes (cont.)

- Assume that the zeros and poles of W(z) are different from each other and that they all lie strictly inside the unit circle
- · Since we are determining the one-step ahead predictor, we get:

$$\begin{array}{c|c}
C(z) \\
A(z) \\
\hline
C(z) - A(z)
\end{array} \quad \begin{vmatrix}
A(z) \\
1
\end{vmatrix}$$

Thus:

$$\frac{C(z)}{A(z)} = 1 + \frac{C(z) - A(z)}{A(z)} = 1 + z^{-1} \frac{z \left[C(z) - A(z)\right]}{A(z)}$$

and hence

$$\widehat{W}_1(z) = \frac{z \left[C(z) - A(z)\right]}{A(z)}$$
$$W_1(z) = \frac{z \left[C(z) - A(z)\right]}{C(z)}$$

One-step Ahead Prediction for ARMA Processes (cont.)

• Since A(z) and C(z) are monic, in C(z) - A(z) the constant term is missing:

$$C(z) - A(z) = (1 + c_1 z^{-1} + \dots + c_n z^{-n}) - (1 - a_1 z^{-1} - \dots - a_n z^{-n})$$

= $(c_1 + a_1) z^{-1} + \dots + (c_n + a_n) z^{-n}$

Hence:

$$C(z) \hat{v}(t+1|t) = [C(z) - A(z)] z v(t)$$

= $[C(z) - A(z)] v(t+1)$
= $[(c_1 + a_1)z^{-1} + \dots + (c_n + a_n)z^{-n}] v(t+1)$
= $(c_1 + a_1)v(t) + (c_2 + a_2)v(t-1) + \dots + (c_n + a_n)v(t-n+1)$

and then:

$$\hat{v}(t+1|t) = -c_1\hat{v}(t|t-1) - c_2\hat{v}(t-1|t-2)\cdots - c_n\hat{v}(t-n+1|t-n) + (c_1+a_1)v(t) + (c_2+a_2)v(t-1) + \cdots + (c_n+a_n)v(t-n+1)$$

Remark. Stability of the predictor guaranteed because zeros of C(z) are assumed to lie inside the unit circle

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One-step Ahead Prediction for ARMA Processes (cont.)

Alternative Procedure

$$A(z)v(t) = C(z)\xi(t)$$

• Add and subtract to the right-hand side the term C(z)v(t):

$$\begin{aligned} A(z)v(t) &= C(z)\xi(t) + C(z)v(t) - C(z)v(t) \\ \implies C(z)v(t) &= [C(z) - A(z)]v(t) + C(z)\xi(t) \\ \implies v(t) &= \frac{[C(z) - A(z)]}{C(z)}v(t) + \xi(t) \quad (\star) \end{aligned}$$

- But $\frac{[C(z) A(z)]}{C(z)} = \#z^{-1} + \#z^{-2} + \cdots$ and hence v(t) in (\star) is a function of $v(t-1), v(t-2), \ldots$
- Moreover $\xi(t)$ is uncorrelated with the past of v(t). Then:

$$\hat{v}(t \,|\, t-1) = \frac{[C(z) - A(z)]}{C(z)} \, v(t)$$

where $\xi(t)$ has been dropped since it is uncorrelated with the first term and it is unpredictable from the past

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Solution of the Prediction Problem

Prediction in Presence of External Inputs

Prediction in Presence of External Inputs

• First, consider the simple case

 $v(t) = av(t-1) + u + \xi(t), \quad |a| < 1, \quad \xi(\cdot) \sim WN(0, \lambda^2)$

where *u* is constant, known, and deterministic.

• Clearly:

$$E[v(t)] = aE[v(t-1)] + u + E[\xi(t)]$$

$$\implies (1-a)E[v(t)] = u \implies E[v(t)] = \frac{u}{1-a}$$

• Set
$$\bar{v} = \frac{u}{1-a}$$
 and $\tilde{v}(t) = v(t) - \bar{v}$. Then:

$$\begin{split} \tilde{v}(t) &= v(t) - \bar{v} = av(t-1) + u + \xi(t) - \bar{v} \\ \implies \tilde{v}(t) &= av(t-1) - a\bar{v} + u + \xi(t) + (a-1)\bar{v} \\ &= a\tilde{v}(t-1) + u + \xi(t) + (a-1)\bar{v} \\ &= a\tilde{v}(t-1) + \xi(t) \end{split}$$

Prediction in Presence of External Inputs (cont.)

· Let us write the process in terms of "variations":

$$\tilde{v}(t) = a\tilde{v}(t-1) + \xi(t)$$

This process is AR(1) and hence:

$$\hat{\tilde{v}}(t \,|\, t-1) = a \tilde{v}(t-1)$$

But $v(t) = \tilde{v}(t) + \bar{v}$ and thus:

$$\hat{v}(t \mid t-1) = \hat{\tilde{v}}(t \mid t-1) + \bar{v} = a\tilde{v}(t-1) + \bar{v} = a[v(t-1) - \bar{v}] + \bar{v} = av(t-1) + u$$

To sum-up:

the one-step ahead predictor can be obtained by adding the known external input to the predictor obtained without considering the external input

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• Let us generalize (without proof) to the case of ARMAX models:

$$A(z)v(t) = B(z)u(t) + C(z)\xi(t)$$

with:

$$A(z) = 1 - a_1 z^{-1} - \dots - a_n z^{-n}$$

$$B(z) = b_1 z^{-1} + \dots + b_n z^{-n}$$

$$C(z) = 1 + c_1 z^{-1} + \dots + c_n z^{-n}$$

The one-step ahead predictor can be obtained by adding the known (deterministic or not) external term B(z)u(t) to the predictor obtained without considering the external input:

$$\hat{v}(t+1|t) = -c_1\hat{v}(t|t-1) - c_2\hat{v}(t-1|t-2)\cdots - c_n\hat{v}(t-n+1|t-n) + (c_1+a_1)v(t) + (c_2+a_2)v(t-1) + \cdots + (c_n+a_n)v(t-n+1) + b_1u(t) + b_2u(t-1) + \cdots + b_nu(t-n+1)$$

Models and Predictors

Consider the general model

$$\mathcal{M}(\vartheta): \quad y(t) = G(z) \, u(t-1) + W(z) \, \xi(t)$$

where ϑ denotes a **vector of parameters characterizing the model** in which the one-step delay between input and output is explicitly enhanced (a widely used convention)



Models and Predictors (cont.)

Let us determine the optimal predictor:

$$\begin{split} y(t) &= G(z) \, u(t-1) + W(z) \, \xi(t) \\ \implies \frac{1}{W(z)} \, y(t) &= \frac{G(z)}{W(z)} \, u(t-1) + \xi(t) \\ \implies y(t) + \frac{1}{W(z)} \, y(t) &= y(t) + \frac{G(z)}{W(z)} \, u(t-1) + \xi(t) \\ \implies y(t) &= \left[1 - \frac{1}{W(z)}\right] \, y(t) + \frac{G(z)}{W(z)} \, u(t-1) + \xi(t) \end{split}$$

• But W(z) is monic and hence $1 - \frac{1}{W(z)} = \#z^{-1} + \#z^{-2} + \cdots$. Therefore, $\left[1 - \frac{1}{W(z)}\right] y(t)$ depends on $y(t-1), y(t-2), \ldots$ • Moreover, $\frac{G(z)}{W(z)} u(t-1)$ depends on $u(t-1), u(t-2), \ldots$

Models and Predictors (cont.)

• Therefore, since $\xi(t)$ is white, the class of optimal predictors $\widehat{\mathcal{M}}(\vartheta)$ associated with the class of models $\mathcal{M}(\vartheta)$ is:

$$\widehat{\mathcal{M}}(\vartheta): \quad \widehat{y}(t \mid t-1) = \left[1 - \frac{1}{W(z)}\right] y(t) + \frac{G(z)}{W(z)} u(t-1)$$

where the optimality stems from the fact that the prediction error

$$\hat{\varepsilon}(t) = y(t) - \hat{y}(t \mid t - 1) = \xi(t)$$

is white (zero expected value and variance equal to the variance of $\xi(t)$).

• Let us now consider another predictor $\widetilde{\mathcal{M}}(\vartheta)$ with a white prediction error $\tilde{\varepsilon}(t)$ with zero expected value. Assume that $\widetilde{\mathcal{M}}(\vartheta)$ is "better" than $\widehat{\mathcal{M}}(\vartheta)$, that is

$$\operatorname{var}\left[\tilde{\varepsilon}(t)\right] < \operatorname{var}\left[\hat{\varepsilon}(t)\right]$$

Models and Predictors (cont.)

• But:

$$\begin{split} \tilde{\varepsilon}(t) &= y(t) - \tilde{y}(t \mid t-1) = y(t) - \hat{y}(t \mid t-1) + \hat{y}(t \mid t-1) - \tilde{y}(t \mid t-1) \\ &= \xi(t) + \hat{y}(t \mid t-1) - \tilde{y}(t \mid t-1) \end{split}$$

On the other hand, $\widehat{\mathcal{M}}(\vartheta)$ and $\widetilde{\mathcal{M}}(\vartheta)$ are predictors and hence:

- $\hat{y}(t \mid t-1)$ depends on $y(t-1), y(t-2), \ldots$
- $\tilde{y}(t \mid t-1)$ depends on $y(t-1), y(t-2), \ldots$

Therefore $\,\hat{y}(t\,|\,t-1)-\tilde{y}(t\,|\,t-1)\,$ is uncorrelated with $\,\xi(t)\,$ and hence

$$\begin{aligned} \operatorname{var}[\tilde{\varepsilon}(t)] &= \operatorname{var}[\xi(t) + \hat{y}(t \mid t-1) - \tilde{y}(t \mid t-1)] \\ &= \operatorname{var}[\xi(t)] + \operatorname{var}[\hat{y}(t \mid t-1) - \tilde{y}(t \mid t-1)] \\ &\geq \operatorname{var}[\xi(t)] = \operatorname{var}[\hat{\varepsilon}(t)] \end{aligned}$$

which **contradicts** the assumption $\operatorname{var} \left[\tilde{\varepsilon}(t) \right] < \operatorname{var} \left[\hat{\varepsilon}(t) \right]$ hence proving that $\widehat{\mathcal{M}}(\vartheta)$ is optimal.

Summing up:

The model and its associated predictor

$$\mathcal{M}(\vartheta): \quad y(t) = G(z) u(t-1) + W(z) \xi(t)$$
$$\implies \quad \widehat{\mathcal{M}}(\vartheta): \quad \widehat{y}(t \mid t-1) = \left[1 - \frac{1}{W(z)}\right] y(t) + \frac{G(z)}{W(z)} u(t-1)$$

 $\widehat{\mathcal{M}}(\vartheta)$ is called **model in prediction form**.

Models and Predictors

Predictors for ARX Models

Predictors for ARX Models

$$\mathcal{M}(\vartheta): \quad A(z) y(t) = B(z) u(t-1) + \xi(t)$$
$$\implies \quad G(z) = \frac{B(z)}{A(z)} \quad W(z) = \frac{1}{A(z)} \qquad \vartheta = \begin{bmatrix} a_1 \\ \vdots \\ a_n \\ b_1 \\ \vdots \\ b_n \end{bmatrix}$$

Then:

$$\hat{y}(t \mid t-1) = \left[1 - \frac{1}{W(z)}\right] y(t) + \frac{G(z)}{W(z)} u(t-1)$$

= $[1 - A(z)] y(t) + B(z) u(t-1)$
= $a_1 y(t-1) + a_2 y(t-2) + \dots + a_n y(t-n)$
+ $b_1 u(t-1) + b_2 u(t-2) + \dots + b_n u(t-n)$

Observe that $\hat{y}(t | t - 1)$ does not depend on its past values, that is, the predictor is not dynamic and hence it is always stable

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Models and Predictors

Predictors for ARMAX Models

Predictors for ARMAX Models

$$\begin{aligned} \mathcal{M}(\vartheta) : \quad A(z) \, y(t) &= B(z) \, u(t-1) + C(z) \, \xi(t) \\ \implies & G(z) = \frac{B(z)}{A(z)} \qquad W(z) = \frac{C(z)}{A(z)} \qquad \vartheta = \begin{bmatrix} a_1 \\ \vdots \\ a_n \\ b_1 \\ \vdots \\ b_n \\ c_1 \\ \vdots \\ c_n \end{bmatrix} \\ \end{aligned}$$
 Then:
$$\hat{y}(t \,|\, t-1) &= \begin{bmatrix} 1 - \frac{1}{W(z)} \end{bmatrix} \, y(t) + \frac{G(z)}{W(z)} \, u(t-1) \\ &= \begin{bmatrix} 1 - \frac{A(z)}{C(z)} \end{bmatrix} \, y(t) + \frac{B(z)}{C(z)} \, u(t-1) \\ &= \begin{bmatrix} C(z) - A(z) \\ C(z) \end{bmatrix} \, y(t) + \frac{B(z)}{C(z)} \, u(t-1) \end{aligned}$$

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Hence:

$$\hat{y}(t | t-1) = -c_1 \hat{y}(t-1 | t-2) - c_2 \hat{y}(t-2 | t-3) \cdots - c_n \hat{y}(t-n | t-n-1) + (c_1 + a_1)y(t-1) + (c_2 + a_2)y(t-2) + \cdots + (c_n + a_n)y(t-n) + b_1 u(t-1) + b_2 u(t-2) + \cdots + b_n u(t-n)$$

Observe that $\hat{y}(t|t-1)$ now depends on its past values, that is, **the** predictor is dynamic.

Therefore, its stability depends on the position in the complex plane of the zeroes of ${\cal C}(z)$

Models and Predictors

Predictors for MA Models

Predictors for MA Models

$$\begin{aligned} \mathcal{M}(\vartheta) : \quad y(t) &= C(z)\,\xi(t) \\ \implies \quad G(z) &= 0 \qquad W(z) &= C(z) \qquad \vartheta &= \left[\begin{array}{c} c_1 \\ \vdots \\ c_n \end{array} \right] \end{aligned}$$

Then:

$$\hat{y}(t \mid t-1) = \left[1 - \frac{1}{W(z)}\right] y(t) + \frac{G(z)}{W(z)} u(t-1) \\
= \left[1 - \frac{1}{C(z)}\right] y(t) = \left[\frac{C(z) - 1}{C(z)}\right] y(t) \\
= -c_1 \hat{y}(t-1 \mid t-2) - c_2 \hat{y}(t-2 \mid t-3) \cdots - c_n \hat{y}(t-n \mid t-n-1) \\
+ c_1 y(t-1) + c_2 y(t-2) + \cdots + c_n y(t-n)$$

Analogously to the ARMAX case, observe that $\hat{y}(t | t - 1)$ depends on its past values, that is, **the predictor is dynamic**.

Therefore, its stability depends on the position in the complex plane of the zeroes of ${\cal C}(z)$.

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Models and Predictors

Predictors for ARXAR Models

Predictors for ARXAR Models

$$\begin{aligned} \mathcal{M}(\vartheta) : \quad A(z) \, y(t) &= B(z) \, u(t-1) + \frac{1}{D(z)} \xi(t) \\ \implies \quad G(z) &= \frac{B(z)}{A(z)} \qquad W(z) = \frac{1}{A(z)D(z)} \qquad \vartheta = \begin{bmatrix} a_1 \\ \vdots \\ a_n \\ b_1 \\ \vdots \\ b_n \\ d_1 \\ \vdots \\ d_n \end{bmatrix} \\ \end{aligned}$$
 Then:
$$\hat{y}(t \,|\, t-1) &= \begin{bmatrix} 1 - \frac{1}{W(z)} \end{bmatrix} y(t) + \frac{G(z)}{W(z)} \, u(t-1) \\ &= \begin{bmatrix} 1 - A(z)D(z) \end{bmatrix} y(t) + B(z)D(z) \, u(t-1) \end{aligned}$$

Analogously to the ARX case, **the predictor is not dynamic and hence it is always stable**

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Models and Predictors

Concluding Remarks

- All models in prediction form $\widehat{\mathcal{M}}(\vartheta)$ depend linearly on y(t) and u(t)
- In general, the stability of the model in prediction form $\widehat{\mathcal{M}}(\vartheta)$ has **nothing to do** with the stability of the associated model $\mathcal{M}(\vartheta)$: for all considered models, the stability depends on the zeroes of A(z) (**poles of the model**) whereas, for the models in prediction form $\widehat{\mathcal{M}}(\vartheta)$, the stability depends on the zeroes of C(z) (**poles of the model in prediction form**)
- Consider the ARX model in prediction form:

 $\hat{y}(t \mid t-1) = a_1 y(t-1) + a_2 y(t-1) + \dots + a_n y(t-n)$ $+ b_1 u(t-1) + b_2 u(t-1) + \dots + b_n u(t-n)$

Hence, $\hat{y}(t | t - 1)$ depends linearly on the parameters a_i, b_i . This property is typically exploited in the identification algorithms • Consider the ARXAR model in prediction form:

$$\hat{y}(t \mid t - 1) = [1 - A(z)D(z)] y(t) + B(z)D(z) u(t - 1)$$

Hence:

- + For a given D(z) , $\hat{y}(t\,|\,t-1)$ depends linearly on the parameters a_i, b_i
- For given A(z), B(z), $\hat{y}(t \mid t-1)$ depends linearly on the parameters d_i

This property is typically exploited in the identification algorithms

Models and Predictors: Remarks (cont.)

• On the other hand, consider a first-order ARMAX model in prediction form:

$$\hat{y}(t \mid t-1) = \left[\frac{C(z) - A(z)}{C(z)}\right] y(t) + \frac{B(z)}{C(z)} u(t-1)$$

But:

$$\left[\frac{C(z) - A(z)}{C(z)}\right] = \frac{(a+c)z^{-1}}{1+cz^{-1}} = (a+c)z^{-1} - c(a+c)z^{-2} + \cdots$$

$$\frac{B(z)}{C(z)} = \frac{b}{1 + cz^{-1}} = b - cbz^{-1} + \cdots$$

Hence, $\hat{y}(t | t - 1)$ depends in a **nonlinear** on the parameters a_i, b_i, c_i .

This nonlinear dependence will make the identification algorithms much more complicated

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Lecture 10 Solution of the Prediction Problem

END